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Realistic Tables with Competing Risks Application to Inception Rates Estimation for Long Term Care Insurance

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Agenda

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2 Modelisation

3 Estimation methods for competing risks

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Introduction		
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Long Term Care Insurance Framework

- In France, LTC insurance provides benefits for elderly people suffering from a loss of mobility and autonomy in their activity of daily living.
- In addition to the social benefits.
- LTC insurance may be individual or collective.
- Payment of benefits depends to the level of dependency.

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Long Term Care Insurance Framework

- Pricing, reserving and managing LTC risks strongly depend to the tables selected.
- LTC risks are tricky to estimate:
 - the first french products dates from the early of 80's,
 - only a few insurers have reliable data with sufficient amount (but for higher levels of dependency and not at older ages),
 - definitions and grids are not unique,
 - covariate effects (gender, place, pathology, etc.) are rarely taking into account.
- Forecasting LTC risks (longevity, disability and mortality) are also a very difficult exercice.
- Pratictioners often use empirical methods and expert opinions.

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Introduction		

Motivation

- Constructing realistic tables for inception rates in dependency is an important challenge in Solvency II perspective.
- Distinguing entry by pathology is very usefull as:
 - waiting periods in contracts depend on the type of desease,
 - pathologies have a major role in the survival of LTC claimants.
- In presence of competing risks, practioners often use techniques based on latent failure times and arbitrary choices for modelling dependence between them.
- Aims of our approach:
 - use a more relevant multistate approach to estimate inception rates with right censored data,
 - measure the biais that the common approach used by prationers comprises.
- Limits of our approach:
 - a non-parametric approach requires a signicant amount of reliable data and does not permit forecasting to older ages,
 - longitudinal data are required.

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Competing Risks Model

• A competing risks process models adequately both the entry in dependency according to different pathologies and other exit causes (e.g. death, cancellation).



Example of exit causes with 4 types of pathology.

	Exit causes
e_1	Neurologic pathologies
e_2	Various pathologies
e_3	Terminal cancers
e_4	Dementia
e_5	Death
<i>e</i> ₆	Cancel

Benefits and premiums are paid according to the papern of states of the policyholder.

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Model		

Quantities of Interest

We introduce a Markov process $(X_t)_{t\geq 0}$ with finite state space $S = \{e_0, e_1, \ldots, e_n\}$ where the state e_j represents the *j*-th exit cause of the initial state. *T* represents the survival time in initial state and *C* the right-censoring time.

Transition probability

$$p_{0j}(s,t) = \mathbb{P} \left(X_t = e_j \mid X_s = e_0 \right)$$

= $\mathbb{P} \left(T \le t, X_T = e_j \mid T > s \right).$

Cause-specific hazard

$$\mu_{0j}(t) = \lim_{\Delta t \to 0} \frac{p_{0j}(t, t + \Delta t)}{\Delta t}$$

Latent failure time

$$T_{0j} = \inf_{t \ge 0} \left(X_t = e_j \right).$$

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Technical provisions

- Technical provisions correspond to the expectation of future discounted cash-flows relating to the contract.
- With a policyholder aged x years at t_0 , reserves at time $t \ge t_0$ are

$$\sum_{j \neq e_0} \int_{t-t_0}^{\infty} B(t, t_0 + \tau) p_{00}(x, x + \tau) \mu_{0j}(x + \tau) c_j(x + \tau) d\tau - \int_{t-t_0}^{\infty} B(t, t_0 + \tau) p_{00}(x, x + \tau) b(x + \tau) d\tau.$$

• Inception rate $q_i(t)$ appears when we approximate the above formula

$$\approx \sum_{j \neq e_0} \sum_{k=t-t_0}^{\infty} B(t, t_0 + k + 1) p_{00}(x, x+k) q_j(x+k) c_j(x+k)$$
$$- \sum_{k=t-t_0}^{\infty} B(t, t_0 + k) p_{00}(x, x+k) b(x+k).$$

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Classical Methods

- A extensive literature is dedicated to infer competing risks model:
 - methods based on latent failure times (e.g. Prentice et al. (1978)),
 - proportional hazards models (e.g. Fine and Gray (1999)),
 - multistate approachs (e.g. Andersen et al. (2002)).
- Practioners often use latent failure times approachs (e.g. Deléglise *et al.* (2009)) but:
 - latent failure time variable T_{0j} is artificial,
 - dependence structure between latent failure times is unknow.
- These last estimators may overestimate the inception rates (see Gooley *et al.* (1999)).

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Multistate Approach

We aim to estimate inception rates with a non-parametric multistate approach (see Andersen *et al.* (1993)).

- We observe continuously *M* independent policyholders $(\bar{T}^m, V^m)_{m=1,...,M}$ where $\bar{T}^m = \min(T^m, C^m)$ and $V = X_T$ the failure cause.
- The cumulative cause-specific hazards are estimated with Nelson-Aalen estimator.
- The inception rates are consequently obtained with the Aalen-Johansen estimator

$$\hat{q}_{j}(t) = \sum_{\left\{m, t < \bar{T}^{(m)} \le t+1\right\}} \hat{S}\left(\bar{T}^{(m)}-\right) \frac{\mathbb{I}\left\{V^{(m)}=j\right\}}{Y_{0}\left(\bar{T}^{(m)}\right)}.$$

• These estimators are asymptotically normally distributed.

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• Pratictioners usually assume latent failure times are independent and then we have

$$\mathbb{P}\left(T_{0j}>t\right)=\exp\left(-\int_{0}^{t}\mu_{0j}\left(\tau\right)d\tau\right).$$

- Consequently, the inception rates per event are estimated with Kaplan-Meier estimator and noted $q_{i}^{*}(t)$.
- Then, arbitrary dependence rule are applied on $q_i^*(t)$ to verify the egality

$$1 - \sum_{j=1}^{n} q_{j}(t) = p_{00}(t, t+1).$$

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• Inception rates adjustment for some order (j_1, \ldots, j_n) is:

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	Estimation	
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- Inception rates adjustment for some order (j_1, \ldots, j_n) is:
 - $\breve{q}_{j_1}(t) = \hat{q}_{j_1}^*(t)$

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- Inception rates adjustment for some order (j_1, \ldots, j_n) is:
 - $\check{q}_{j_1}(t) = \hat{q}_{j_1}^*(t)$
 - $\breve{q}_{j_2}(t) = \hat{q}_{j_2}^*(t) \left(1 \hat{q}_{j_1}^*(t)\right),$

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- Inception rates adjustment for some order (j_1, \ldots, j_n) is:
 - $\breve{q}_{j_1}(t) = \hat{q}_{j_1}^*(t)$
 - $\breve{q}_{j_2}(t) = \hat{q}_{j_2}^*(t) \left(1 \hat{q}_{j_1}^*(t)\right),$
 - ...

•
$$\breve{q}_{j_n}(t) = \hat{q}_{j_n}^*(t) \prod_{k=1}^{n-1} (1 - \hat{q}_{j_k}^*(t)).$$

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- Inception rates adjustment for some order (j_1, \ldots, j_n) is:
 - $\breve{q}_{j_1}(t) = \hat{q}_{j_1}^*(t)$
 - $\breve{q}_{j_2}(t) = \hat{q}_{j_2}^*(t) \left(1 \hat{q}_{j_1}^*(t)\right),$
 - ... • $\breve{q}_{j_n}(t) = \hat{q}_{j_n}^*(t) \prod_{k=1}^{n-1} (1 - \hat{q}_{j_k}^*(t)).$
- The new estimator $\breve{q}_{j}(t)$ has upper and lower bounds

$$b_{j}^{-}(t) = \hat{q}_{j}^{*}(t) \prod_{k \neq j} (1 - \hat{q}_{k}^{*}(t)) \leq \breve{q}_{j}(t) \leq \hat{q}_{j}^{*}(t) = b_{j}^{+}(t).$$

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Estimation with Multistate Approach



Figure: Inception rates estimate with approximate pointwise 95% confidence intervals

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Smoothed Rates

- We smooth the crude inception rates with non-parametric Whittaker-Henderson model.
- Smoothing parameters are selected on the basis of residuals analysis and by regarding the following fitting criteria:
 - cross validation and generalized cross validation,
 - AIC and AICC.

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Smoothed Rates



Figure: Smoothed rates, residuals and Student residuals

Introduction	Model 000	Estimation 0000	Numerical Application	Summary
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Measuring Estimation Risk

- We consider systematic risks induced by the construction of such tables and resulting from:
 - estimation of crude rates (due to sampling variation),
 - parameters estimation.
- These risks are taken into account with non-parametric bootstrap.

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Measuring Estimation Risk



Figure: Simulated rates (K = 1000) with 95% simulated confidence intervals

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Measuring Estimation Risk

We compute dispersion coefficients

$$c\left(\psi_{jx}\right) = \frac{\sqrt{\sum_{k=1}^{K} \left(\tilde{q}_{j}^{k}\left(x\right) - \tilde{q}_{j}\left(x\right)\right)^{2}}}{\tilde{q}_{j}\left(x\right)}$$

• The risk is relatively important and should be considered carefully for technical provisions valuation.

Exit causes	Average $c(\psi_{jx})$
Neurologic pathologies	6.01%
Various pathologies	12.12%
Terminal cancers	7.94%
Dementia	6.53%
Death	2.05%
Cancel	3.76%

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Comparing with the latent failure times approach



Figure: Comparing multistate and latent failure times approachs

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Comparing with the latent failure times approach

- We compare 4 differents priority orders.
- At 65 years old, the larger gap measured on technical provisions is around 2.5%.
- Priority orders should be selected soundly!



Figure: Gap on technical provisions with the both approachs

		Summary

Summary

- Multistate approachs are rarely used by pratictioners which prefer Kaplan-Meier estimators and marginal approachs for crude rates estimation.
- Multistate approachs are a little more complex but theorically more appropriate for competing risks as the joint distribution of latent failure times are unobserved.
- A non-parametric approach may difficult to implement due to the lack of available data but:
 - it provides an realistic fit that one can expected in Solvency II perspective.
 - it can be used to perform goodness-of-fit tests.
- Biais observed with latent failure times approach depends on the treatment of priorities applied on each cause. In our application, a sound choice may reduce significatively this biais and justify the use of latent failure times approach.
- Outlook
 - Extrapoling inception rates to older ages.
 - Taking into account other sources of heterogeneity.

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		Summary

Thank you for your kind attention.

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